

- 1.39 a. Suppose  $A$  and  $B$  are mutually exclusive. Then  $A \cap B = \emptyset$  and  $P(A \cap B) = 0$ . If  $A$  and  $B$  are independent, then  $0 = P(A \cap B) = P(A)P(B)$ . But this cannot be since  $P(A) > 0$  and  $P(B) > 0$ . Thus  $A$  and  $B$  cannot be independent.
- b. If  $A$  and  $B$  are independent and both have positive probability, then

$$0 < P(A)P(B) = P(A \cap B).$$

This implies  $A \cap B \neq \emptyset$ , that is,  $A$  and  $B$  are not mutually exclusive.

1.49 For every  $t$ ,  $F_X(t) \leq F_Y(t)$ . Thus we have

$$P(X > t) = 1 - P(X \leq t) = 1 - F_X(t) \geq 1 - F_Y(t) = 1 - P(Y \leq t) = P(Y > t).$$

And for some  $t^*$ ,  $F_X(t^*) < F_Y(t^*)$ . Then we have that

$$P(X > t^*) = 1 - P(X \leq t^*) = 1 - F_X(t^*) > 1 - F_Y(t^*) = 1 - P(Y \leq t^*) = P(Y > t^*).$$

2.20 From Example 1.5.4, if  $X$  = number of children until the first daughter, then

$$P(X = k) = (1 - p)^{k-1}p,$$

where  $p$  = probability of a daughter. Thus  $X$  is a geometric random variable, and

$$\begin{aligned} EX &= \sum_{k=1}^{\infty} k(1-p)^{k-1}p = p - \sum_{k=1}^{\infty} \frac{d}{dp}(1-p)^k = -p \frac{d}{dp} \left[ \sum_{k=0}^{\infty} (1-p)^k - 1 \right] \\ &= -p \frac{d}{dp} \left[ \frac{1}{p} - 1 \right] = \frac{1}{p}. \end{aligned}$$

Therefore, if  $p = \frac{1}{2}$ , the expected number of children is two.