

The gradient algorithm for parameter and output estimation for dual-rate CARARMA systems \star

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Abstract: A recursive generalized extended stochastic forgetting gradient algorithm is used to identify the dual-rate stochastic systems based on the polynomial transformation technique. A time-varying forgetting factor is included to improve the rate of convergence. The intersample output estimation algorithm is also studied in the paper. Finally, a simulation example shows that the algorithm is excellent effective in parameter identification and output estimation.

1. INTRODUCTION

This paper deals with a dual-rate systems where the output y(kJT) is sampled at a slower rate than the input u(kT) and the output is needed at the fast rate. The least squares algorithms are used to estimate the parameters for dual-rate systems (Gudi et al. 1995; Huang et al. 1993; Lu and Fisher 1988,1989; Ding and Chen 2005a). But these algorithms mainly deal with the noise-free systems. In multirate stochastic systems with noise, Ding and Chen al. have proposed least squares algorithms for dual-rate systems and intersample estimation for the dual-rate CAR model, especially the convergence properties of the algorithms (Ding and Chen 2004a,2004b,2004c,2005b). Ding and Chen (2002,2007) have also studied the convergence properties of the stochastic gradient algorithm for CAR and CARMA models in single-rate system. Compared with the least squares, the forgetting gradient algorithm has low computation burden and the convergence rate of the parameter estimation can improve due to the presence of forgetting factor λ . So Yang and Tian (2006)has studied the dual-rate stochastic forgetting gradient algorithm for the CAR model. But for more complicated dual-rate CARARMA model, parameter identification and output estimation haven't been studied so for. The main purpose of this paper is parameter identification and output estimation for the dual-rate CARARMA system.

2. THE DESCRIPTION OF THE ALGORITHM

Assume that linear discrete time-invariant single-rate process model is represented by the following equation,

$$a(z)y(kT) = b(z)u(kT), \tag{1}$$

$$a(z) = 1 + a_1 z^{-1} + a_2 z^{-2} + \dots + a_n z^{-n} = \prod_{i=1}^n [1 - \lambda_i z^{-1}],$$

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$$b(z) = b_1 z^{-1} + b_2 z^{-2} + \dots + b_n z^{-n}.$$

u(kT) and y(kT) are the input and output sequences of the system, respectively. The available measurement data are u(kT) and y(kJT), and the [y(kJT + iT), i = 1, ..., J - 1]are not available due to the sensor constraints in output sampling. Then the polynomial transformation technique in (Ding and Chen 2005c) is used to transform the model (1) into a form which can be identified from the available dual-rate data.

Define the following polynomial equation

$$\phi_J(z) = \prod_{i=1}^n (1 + \lambda_i z^{-1} + \lambda_i^2 z^{-2} + \dots + \lambda_i^{J-1} z^{-J+1}).$$

Multiplying both sides of equation(1) by $\phi_I(z)$ gets a new model: $\alpha(z^J)y(kJ) = \beta(z)u(kJ),$

where

$$\alpha(z^{J}) = \phi_{J}(z)a(z) = 1 + \alpha_{1}z^{-J} + \alpha_{2}z^{-2J} + \dots + \alpha_{n}z^{-nJ},$$

 $\beta(z) = \phi_J(z)b(z) = \beta_1 z^{-1} + \beta_2 z^{-2} + \dots + \beta_n J z^{-nJ}.$ Introduce Autoregressive Moving average noise model Equation (2) can be expressed as

$$\alpha(z^J)y(kJ) = \beta(z)u(kJ) + \frac{d(z)}{c(z)}v(kJ).$$
(3)

where

$$c(z) = 1 + c_1 z^{-1} + c_2 z^{-2} + \dots + c_{n_c} z^{-n_c},$$

$$l(z) = 1 + d_1 z^{-1} + d_2 z^{-2} + \dots + d_{n_d} z^{-n_d}.$$

Defining

$$e(kJ) = \frac{d(z)}{c(z)}v(kJ).$$
(4)

(2)

Equation (3) and (4) can be written as the recursive equation,

$$y(kJ) = \varphi_s^T(kJ)\theta_s + e(kJ), \tag{5}$$

$$e(kJ) = \varphi_n^T(kJ)\theta_n + v(kJ).$$
(6)

where

$$\theta_s^T(kJ) := [\alpha_1(kJ), \alpha_2(kJ), \dots, \alpha_n(kJ); \beta_1(kJ), \beta_2(kJ), \dots, \beta_{nJ}(kJ)],$$

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$$\begin{split} \theta_n^T(kJ) &:= [c_1(kJ), c_2(kJ), ..., c_{n_c}(kJ); \\ & d_1(kJ), d_2(kJ), ..., d_{n_d}(kJ)], \\ \varphi_s^T &:= [-y(kJ-J), -y(kJ-2J)..., -y(kJ-nJ); \\ & u(kJ-1), u(kJ-2), ..., u(kJ-nJ)], \\ \varphi_n^T &:= [-e(kJ-1), -e(kJ-2)..., -e(kJ-n_c); \\ & v(kJ-1), v(kJ-2), ..., v(kJ-n_d)]. \end{split}$$

To simplify the computation process, (5) and (6) can be combined into the following equation,

$$y(kJ) = \varphi^T(kJ)\theta + v(kJ),$$

where

$$\begin{aligned} \theta^T(kJ) &= [\theta^T_s(kJ), \theta^T_n(kJ)], \\ \varphi^T(kJ) &= [\varphi^T_s(kJ), \varphi^T_n(kJ)]. \end{aligned}$$

The recursive extended identification algorithm can be used to identify this dual-rate CARARMA model by increasing the dimensions of the parameter vector and information vector, adding the parameters c_i and d_i of the noise model to the parameter vector and adding the noise terms $e(kJ - i), i = 1, ..., n_c$ and $v(kJ - i), i = 1, ..., n_d$ to the information vector. Since e(kq - i) and v(kq - i) are unknown, they are replaced by the estimation $\hat{e}(kq - i)$ and $\hat{v}(kq - i)$,

$$\hat{e}(kJ-i) = y(kJ-i) - \varphi_s^T(kJ-i)\theta_s(kJ), i = 1, ..., n_c, \ (8)$$

$$\hat{v}(kJ-i) = y(kJ-i) - \varphi^T(kJ-i)\hat{\theta}(kJ), i = 1, ..., n_d,$$
(9) or,

$$\hat{v}(kJ-i) = \hat{e}(kJ-i) - \varphi_n^T(kJ-i)\hat{\theta}_n(kJ), i = 1, ..., n_d.$$

But actually the right-hand sides of (8) and (9) contain
the unknown intersample outputs $y(kJ-i)$. Thus, it is
impossible to compute the $\hat{e}(kJ-i)$ and $\hat{v}(kJ-i)$ by
(8) and (9). A feasible way is that after computing the
parameter estimate $\hat{\theta}$ at each step, $y(kJ-i)$ is replaced
by its estimation $\hat{y}(kJ-i)$. Ding and Chen has given a
method to determine the single-rate model to compute the
 $\hat{y}(kJ-i)$ (Ding and Chen 2005a). But a more simple way
is to compute the $\hat{y}(kJ-i)$ directly from the following
equation.

$$\hat{y}(kJ+i) = \begin{cases} y(kJ) , & i = 0, \\ \hat{\varphi_s}^T(kJ+i)\hat{\theta_s}(kJ), & i = 1, \dots, J-1, \end{cases}$$
(10)

where

$$\hat{\varphi_s}(kJ+i) = [-\hat{y}(kJ-J+i), -\hat{y}(kJ-2J+i), \\ \cdots, -\hat{y}(kJ-nJ+i) \\ u(kJ+i), u(kJ+i-1), \\ \cdots, u(kJ+i-nJ)]^T.$$

Then, the forgetting factor gradient algorithm of estimating the parameter vector of the dual-rate stochastic systems in (3) may be expressed as dual-rate recursive generalized extended forgetting gradient algorithm (DR-RGEFG algorithm for short):

$$\hat{\theta}(kJ) = \hat{\theta}(kJ - J) + \frac{\hat{\varphi}(kJ)}{r(kJ)}$$
$$\times [y(kJ) - \hat{\varphi}^T(kJ)\hat{\theta}(kJ - J)], \qquad (11)$$

$$\hat{\theta}(kJ+i) = \hat{\theta}(kJ), i = 1, \cdots, J-1,$$
 (12)

$$r(kJ) = \lambda r(kJ - J) + \|\hat{\varphi}(kJ)\|^2,$$

$$r(0) = 1, \ 0 \le \lambda \le 1,$$

$$\hat{\theta}^{T}(kJ) = [\hat{\theta}_{s}^{T}(kJ), \hat{\theta}_{n}^{T}(kJ)],$$

$$\hat{\varphi}^{T}(kJ) = [\varphi_{s}^{T}(kJ), \hat{\varphi}_{n}^{T}(kJ)],$$

$$\hat{\theta}^{T}(kJ) = [\varphi_{s}^{T}(kJ), \varphi_{n}^{T}(kJ)],$$

$$(13)$$

$$\begin{split} \theta_s^{T}(kJ) &= [\alpha_1(kJ), \alpha_2(kJ), ..., \alpha_n(kJ)], \\ \beta_1(kJ), \beta_2(kJ), ..., \beta_{nJ}(kJ)], \\ \theta_n^{T}(kJ) &= [c_1(kJ), c_2(kJ), ..., c_{n_c}(kJ); \\ d_1(kJ), d_2(kJ), ..., d_{n_d}(kJ)], \end{split}$$

$$\begin{split} \varphi_s^T &= [-y(kJ-1), -y(kJ-2)..., -y(kJ-n); \\ &u(kJ-1), u(kJ-2), ..., u(kJ-nJ)], \\ \hat{\varphi}_n^T &= [-\hat{e}(kJ-1), -\hat{e}(kJ-2)..., -\hat{e}(kJ-n_c); \\ &\hat{v}(kJ-1), \hat{v}(kJ-2), ..., \hat{v}(kJ-n_d)], \end{split}$$

$$\begin{split} \hat{e}(kJ-i) &= \hat{y}(kJ-i) - \varphi_s^T(kJ-i)\hat{\theta_s}(kJ), i = 1, ..., n_c, \\ \hat{v}(kJ-i) &= \hat{y}(kJ-i) - \hat{\varphi}^T(kJ-i)\hat{\theta}(kJ), i = 1, ..., n_d, \\ \mathbf{r} \end{split}$$

$$\hat{v}(kJ-i) = \hat{e}(kJ-i) - \hat{\varphi}_n^T(kJ-i)\hat{\theta}_n(kJ), i = 1, ..., n_d.$$

3. EXAMPLES

Assume that discrete-time system model takes the following form,

$$p(z) = \frac{b(z)}{a(z)},$$

$$a(z) = 1 + a_1 z^{-1} + a_2 z^{-2} = 1 - 1.5 z^{-1} + 0.7 z^{-2},$$

$$b(z) = b_1 z^{-1} + b_2 z^{-2} = z^{-1} + 0.5 z^{-2}.$$

The corresponding dual-rate models with the noise can be expressed as

$$\alpha(z^J)y(kJ) = \beta(z)u(kJ) + \frac{d(z)}{c(z)}v(kJ),$$

where

(7)

$$\phi_J(z) = 1 + 1.5z^{-1} + 0.7z^{-2},$$

$$c(z) = 1 - z^{-1} + 0.41z^{-2}.$$

$$d(z) = 1 + 0.1z^{-1}.$$

Taking J = 2, i.e., $\{u(k), y(2k)\}$ are available data. Here $\{u(k)\}$ is taken as a persistent excitation sequence with zero mean and unit variance, and $\{v(k)\}$ as a white noise sequence with zero mean and variance σ_v^2 . Applying the DR-RGEFG algorithm to estimate the parameters (α_i, β_i) of this system. The parameter estimates with different data length are shown in Table 1, where δ is the relative parameter estimation error measured in the Euclideannorm: $\delta = \|\hat{\theta}(kJ) - \theta\| / \|\theta\|$. θ is the real parameter value, $\hat{\theta}$ is the estimation of the θ .

Fig. 1 shows δ with different $\lambda's$. The smaller λ is, the faster parameter estimation converges. But a smaller λ results in large parameter estimation variance. Though the bigger λ results in slow convergence rate, the parameter estimation error becomes large. A feasible method is to change λ with different k. In this example, take $\lambda = 0.7 + \frac{0.3*k}{length}$, where length is the data length. When $k = 0, \lambda = 0.7$, and when $k = length, \lambda = 1$. Fig. 2 illustrates parameter estimation verse k and Fig. 3 shows the parameter estimation error δ when $\lambda = 0.7 + \frac{0.3*k}{length}$, it can be seen that the parameter estimation is smooth

and fast. Fig. 4 illustrates the simulation run for output estimation error(OEE) $y(k) - \hat{y}(k)$. y(k) is system output and $\hat{y}(k)$ is the output estimation. In Fig. 3, the OEE becomes smaller when $k \to \infty$.

Table 1. The DR-RGEFG estimate of parameter ($\sigma_v^2=1.00, \delta_{ns}=63.695\%)$

λ	k	α_1	α_2	β_1
	500	-0.63548	0.20294	0.10818
1	1000	-0.66862	0.24597	0.12233
	2000	-0.69826	0.27932	0.13572
	500	-0.94082	0.54884	0.74686
0.9	1000	-0.89598	0.52241	0.93541
	2000	-0.86304	0.50453	0.99454
0.7	500	-0.87209	0.50621	0.99909
	1000	-0.84367	0.49993	1.00664
	2000	-0.84821	0.49346	1.00856
	500	-0.86079	0.49886	1.00348
0.5	1000	-0.83715	0.50045	1.00811
	2000	-0.84771	0.49164	1.01184
$\lambda = 0.7 + \frac{0.3 * k}{lanath}$	500	-0.87447	0.50767	0.99508
	1000	-0.84877	0.49996	1.00629
length	2000	-0.85003	0.49480	1.00210
	θ	-0.85000	0.49000	1.00000

λ	k	β_2	β_3	β_4	δ
	500	0.44479	0.28311	0.14316	0.76142
1	1000	0.46894	0.30332	0.15282	0.74634
	2000	0.49226	0.32334	0.16115	0.73247
	500	1.54987	1.01610	0.21262	0.24352
0.9	1000	1.87374	1.28193	0.23201	0.08941
	2000	1.99018	1.42972	0.31292	0.01674
	500	1.98552	1.40821	0.30059	0.02508
0.7	1000	1.99856	1.45006	0.34654	0.00490
	2000	1.99614	1.45673	0.35669	0.00486
	500	2.00366	1.44614	0.35877	0.00618
0.5	1000	2.00115	1.45636	0.34829	0.00686
	2000	1.99632	1.45479	0.35516	0.00508
	500	1.97891	1.39632	0.28843	0.03136
$\lambda = 0.7$	1000	1.99536	1.44379	0.34140	0.00579
$+\frac{0.3*k}{length}$	2000	1.99887	1.45551	0.35574	0.00335
	θ	2.00000	1.45000	0.35000	



Fig. 1. Parameter estimation $\hat{\theta}$ verses k, $(\sigma_v^2=1.00^2, \delta_{ns}=63.695\%)$

4. CONCLUSIONS

A dual-rate recursive generalized extended forgetting gradient algorithm is used to estimate the parameters of the



Fig. 2. Parameter estimation $\hat{\theta}$ verses k, $(\sigma_v^2 = 1.00^2, \delta_{ns} = 63.695\%)$



Fig. 3. The estimation errors δ verses k with $\lambda = 0.7 + \frac{0.3 * k}{length}, (\sigma_v^2 = 1.00^2, \delta_{ns} = 63.695\%)$





dual-rate CARARMA systems which the input is sampled J times faster than the output; The algorithm uses only dual-rate measurement data. For the unknown noises v(k) and e(k), we replace v(k) and e(k) with $\hat{v}(k)$ and $\hat{e}(k)$ which applies the output estimation $\hat{y}(k)$. The result shows that parameter estimation error is small. Based on the estimated models, intersample output estimation is also studied. The example shows the effectiveness of the algorithm.

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