

# Shooting balloons



- $N$  trial subjects,  $i = 1, 2, \dots, N$
- Each shot  $n_i$  times, trying to hit balloons.
- Count hits  $y_i$ .
- Explanatory variables:
  - ▶ Experienced / non-experienced gunman
  - ▶ Wind speed

# Shooting balloons, data



Trail person	1	2	3	...
Experiences	1	0	0	...
Wind speed	2.13	0.59	1.03	...
$n_i$	6	3	5	...
$y_i$	2	1	1	...

## Generalized Linear Model

- 1 Likelihood;  $f(y; \theta)$ , member of the exponential family.
- 2 Link function;  $g(\mu_i) = x_i\beta$ , where  $\mu_i = E(Y_i)$  and  $g()$  is monotone and differentiable.
- 3 Linear component of explanatory variables;  $g(\mu) = X\beta$

# Shooting balloons, model



- $Y_i \sim \text{bin}(n_i, \pi_i)$ ,  $i = 1, 2, \dots, N$
- $\eta_i = \text{logit}(\pi_i) = \log\left(\frac{p_i}{1-p_i}\right)$
- - 1  $\eta_i = \beta_0 \Rightarrow Y_i \sim \text{bin}(n_1, \pi)$
  - 2  $\eta_i = \beta_0 + \beta_1 x_1$
  - 3  $\eta_i = \beta_0 + \beta_2 x_2$
  - 4  $\eta_i = \beta_0 + \beta_1 x_1 + \beta_2 x_2$

Where  $x_1 = 1$  for experienced gunman, otherwise  $x_1 = 0$  and  $x_2$  is wind speed.

## Exponential family

$f(y; \theta)$  belongs to the exponential family if

$$f(y; \theta) = \exp[a(y)b(\theta) + c(\theta) + d(y)]$$

## Score statistics

Let  $l(\theta; y_i)$  be log-likelihood function. Then the score statistic is:

$$U(\theta; y) = \frac{\partial l(\theta; y_i)}{\partial \theta}$$

## Information

Let  $U = U(\theta; Y)$  be the score statistic. Then the information is

$$\mathfrak{I} = \text{Var}(U)$$

If  $Y_i$  has pdf from exponential family:

- $\mathfrak{I} = E(U^2) = -E\left(\frac{\partial U}{\partial \theta}\right) = -E\left(\frac{\partial^2 l(\theta; y)}{\partial \theta^2}\right)$

# Hight of male students

- In population:  $Y \sim (179.8, 6.5^2)$
- Mean of 91 male students: 181.9
- NTNU students higher then Norwegians?

$\chi^2()$  results ch 1.4 and 1.5

### Definition $\chi^2$

If  $Z \sim N(0, 1)$ , then  $Z^2 \sim \chi^2(1)$ .

If  $Z_1, Z_2, \dots, Z_n$  are independent identical distributed  $Z_i \sim N(0, 1)$ , the  $\sum_{i=1}^n Z_i^2 \sim \chi^2(n)$

### Non iid

If  $Y \sim MVN(\mu, \Sigma)$ , then  $(Y - \mu)^T \Sigma^{-1} (Y - \mu) \sim \chi^2(n)$

### Subtraction

If  $X_1^2 \sim \chi^2(m)$  and  $X_2^2 \sim \chi^2(k)$ ,  $m > k$ , and  $X_1^2$  and  $X_2^2$  are independent, we have:  $X^2 = X_1^2 - X_2^2 \sim \chi^2(m - k)$

## 5.3 Taylor series approximations for log-likelihood

Taylor approximations for  $l(\beta)$  near estimate  $b$ :

$$l(\beta) = l(b) + (\beta - b)U(b) + \frac{1}{2}(\beta - b)^2 U'(b)$$

Approximate  $U'(b)$  with  $E(U') = -\mathfrak{S}(b)$ :

$$l(\beta) = l(b) + (\beta - b)U(b) - \frac{1}{2}(\beta - b)^2 \mathfrak{S}$$

For a vector  $b$

$$l(\beta) = l(b) + (\beta - b)U(b) - \frac{1}{2}(\beta - b)^T \mathfrak{S}(\beta - b)$$



## 5.3 Taylor series approximations for log likelihood

$$U(\beta) = U(b) - (\beta - b)U'(b)$$

Approximate  $U'(b)$  with  $E(U') = -\mathfrak{S}(b)$ :

$$U(\beta) = U(b) - (\beta - b)\mathfrak{S}(b)$$