

Problem 1

a) K-M estimator:
$$\hat{R}(t) = \sum_{t_i \leq t} \frac{n_i - d_i}{n_i}$$

t_i	F_i	n_i	d_i	$\frac{n_i - d_i}{n_i}$	$\hat{R}(t_i)$
27		10	1	$\frac{9}{10}$	0.9 0.9
124		8	1	$\frac{7}{8}$	0.7825
136		7	1	$\frac{6}{7}$	0.625
170		6	1	$\frac{5}{6}$	0.5625
210		4	1	$\frac{3}{4}$	0.42825

$\hat{R}(180) = \hat{R}(170) = 0.5625$

Greenwood's formula:

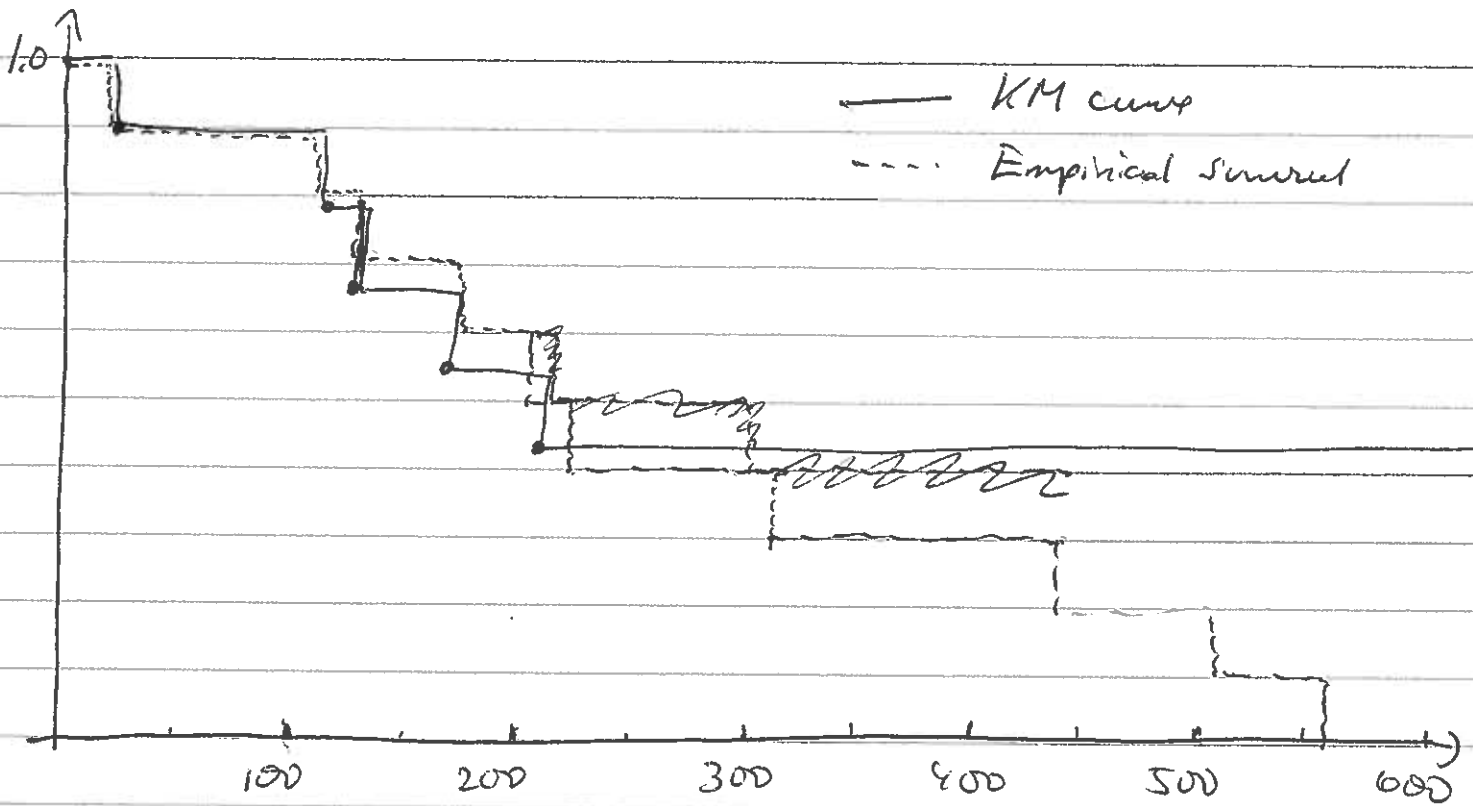
$$\text{Var } \hat{R}(t) = (\hat{R}(t))^2 \sum_{t_i \leq t} \frac{d_i}{n_i(n_i - d_i)}$$

So
$$\text{Var } \hat{R}(180) = \frac{0.5625^2}{\cancel{0.42825}} \left(\frac{1}{10 \cdot 9} + \frac{1}{8 \cdot 7} + \frac{1}{7 \cdot 6} + \frac{1}{6 \cdot 5} \right)$$

$$= \cancel{0.1238} \cdot 0.1651^2$$

So
$$\text{SD } \hat{R}(180) = 0.1651$$

~~$\hat{R}(180)$~~
 ~~$\hat{R}(210)$~~



6) Order the times: 27, 134, 136, 170, 191, 210, 308, 441, 511, 559

$$R^*(180) = \frac{\cancel{X}}{10} = \frac{\cancel{0.6}}{10} = \frac{6}{10}$$

Hence $X \sim \text{bin}(10, R(180))$

$$\text{So } \text{Var}(R^*(180)) = \frac{R(180)(1-R(180))}{10}$$

which is estimated by $\text{Var } R^*(180) = \frac{0.6 \cdot 0.4}{10}$

$$= 0.1549^2$$

$$\text{So } \text{SD } R^*(180) = 0.1549$$

The ~~can~~ It is seen that the estimates are fairly similar, but as is reasonable, the estimated st. dev. of the estimates for $R(t)$ is lower when we have complete data.

c) Know that $E(T) = \int_0^{\infty} R(t) dt$ in general

So ideally one would use the area under the KM-curve from 0 to ∞ to estimate $E(T)$. But this is ∞ , so it could be reasonable to integrate only to the maximum observed time, here 410 for the censored data.

$$\text{Area} = 27 \cdot 1 + 97 \cdot 0.9 + 12 \cdot 0.7825 + 34 \cdot 0.625$$

$$+ 40 \cdot 0.5625 + 200 \cdot 0.42125$$

$$= \underline{\underline{253.525}} \quad \leftarrow \hat{E(T)}$$

From complete data: $E(T) = \bar{T} = \underline{\underline{267.7}}$

Problem 2

a) Curves are ~~$W_j(t)$~~ $N_j^*(t) = \# \text{ failures in } (0, t)$, for $0 < t < 30$

which are unbiased estimators for

$$W_j(t) = E[N_j^*(t)] \quad \text{for } j=1,2.$$

$W_j'(t) = w_j(t)$ so increasing (decreasing) trend is equivalent to convex (concave) plot. There seems to be a convexity here, indicating increasing trends.

$\hat{W}_j(30)$ is expected # failures in a month, for machine j .

$$W_j(t) \sim \text{Poisson}(W_j(t))$$

$$\text{so } \hat{W}_j(30) \sim \text{Poisson}(W_j^*(30))$$

$$\hat{W}_1(30) = 4, \quad \text{SD}(\hat{W}_1(30)) = \sqrt{4} = 2 \quad \text{since } W_j(t) \sim \text{Poisson}.$$

$$\hat{W}_2(30) = 7, \quad \text{SD}(\hat{W}_2(30)) = \sqrt{7} = 2.65$$

b) $H_0: w_j(t)$ is constant vs. $H_1: w_j(t)$ is increasing

$$\text{Test statistics: } Z_j = 2 \sum_{i=1}^{N_j} \ln \frac{c_j}{S_{ij}}$$

$$Z_1 = 2 \left[\ln \frac{30}{11} + \ln \frac{30}{19} + \ln \frac{30}{21} + \ln \frac{30}{28} \right]$$

$$= 3.77$$

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Under H_0 is $Z_1 \sim \chi^2_{2.4=8}$ so reject if

$$\underline{\underline{Z_1}} \leq \chi^2_{0.95, 8} = \underline{\underline{2.73}}$$

Thus: Do not reject H_0 for Machine 1.

$$\begin{aligned} Z_2 &= 2 \left[\ln \frac{30}{8} + \ln \frac{30}{13} + \ln \frac{30}{18} + \ln \frac{30}{21} \right. \\ &\quad \left. + \ln \frac{30}{25} + \ln \frac{30}{26} + \ln \frac{30}{29} \right] \\ &= 6.77 \end{aligned}$$

Under H_0 is $Z_2 \sim \chi^2_{2.7=14}$, so reject if

$$Z_2 \leq \chi^2_{0.95, 14} = 6.57$$

Thus: Do not reject H_0 for Machine 2.

Pooled Mil Hbk test:

H_0 : both $w_1(t)$ and $w_2(t)$ are constant,
but not necessarily equal

vs

H_1 : at least one of them is increasing.

$$Z_{\text{pooled}} = Z_1 + Z_2 = 3.77 + 6.77 = 10.54$$

Under H_0 is $Z_{\text{pooled}} \sim \chi^2_{8+14=22}$, so reject H_0 if

$$Z_{\text{pooled}} \leq \chi^2_{0.975, 22} = 10.98 \text{ for } 0.975$$

Thus we reject H_0 , and conclude that at least one of the $w_j(t)$ is increasing.

Problem 3

a) $N_j \sim \text{Poisson}(W_j(t)) \quad ; j=1,2.$

Here $W_1(t) = \lambda t^\beta \equiv E(N_1)$

$$W_2(t) = \delta \lambda t^\beta \equiv E(N_2)$$

See that $\delta = \frac{E(N_2)}{E(N_1)}$. ~~so δ is the~~

b) General: $L = \left\{ \prod_{i=1}^{N(\tau)} w(S_i) \right\} e^{-W(\tau)}$

so log-likelihood is

$$l = \left\{ \sum_{i=1}^{N(\tau)} \ln w(S_i) \right\} - W(\tau)$$

$$l(\lambda, \beta, \delta) = \sum_{i=1}^{N_1} \ln(\lambda + \ln \beta + (\beta - 1) \ln S_{i1}) - \lambda \cdot 30^\beta$$
$$+ \sum_{i=1}^{N_2} (\ln \delta + \ln \lambda + \ln \beta + (\beta - 1) \ln S_{i2}) - \delta \cdot \lambda \cdot 30^\beta$$

$$= (N_1 + N_2) \ln \lambda + (N_1 + N_2) \ln \beta$$

$$+ N_2 \ln \delta + (\beta - 1) U - \lambda(1 + \delta) \cdot 30^\beta$$

$$\text{where } U = \sum_{i=1}^{N_1} \ln S_{i1} + \sum_{i=1}^{N_2} \ln S_{i2}$$

$$c) \frac{\lambda}{\delta} = \frac{N_2}{N_1} = \frac{7}{4} = 1.75$$

$$\beta = \frac{11}{11 \cdot \ln 30 - 32.1426} = 2.0871$$

$$\lambda = \frac{11}{(1 + 1.75) \cdot 30^{2.0871}} = 0.003305$$

The parameters are all positive, so we use the standard interval for positive parameters given as (general)

$$\theta \pm 1.96 \frac{\sqrt{SE(\hat{\theta})}}{\hat{\theta}}$$

So for λ : $0.003305 \cdot e^{\pm 1.96 \frac{\sqrt{5.278 \cdot 10^{-5}}}{0.003305}}$

$$(4.449 \cdot 10^{-5}, 0.2436)$$

For β : $2.0871 \cdot e^{\pm 1.96}$
 $(1.156, 3.769)$

← Only this needs to be done.

For δ : $(0.5123, 5.978)$

Confint. can be used to test $H_0: \beta = 1$ vs $H_1: \beta \neq 1$ with sign level 5%. Reject H_0 if 1 \notin confint. This is the case, so we reject!

$H_0: \beta = 1$ corresponds to $w_1(t) = 1, w_2(t) = \delta 1$ which is exactly that both $w_1(t)$ and $w_2(t)$ are constant, but possibly different

d) Test statistic $W = 2 [l(\hat{\lambda}^1, \hat{\beta}^1, \hat{\delta}^1) - l(\hat{\lambda}(\beta), \hat{\delta}(\beta))]$

is approx χ^2_2 under H_0 , where $\hat{\lambda}(\beta), \hat{\delta}(\beta)$ are the maximum likelihood estimates of λ and δ when $\beta = 1$.

We need to find $\hat{\lambda}(1), \hat{\delta}(1)$:

When $\beta = 1$ is the log-likelihood

$$l(\lambda, 1, \delta) = (N_1 + N_2) \ln \lambda + N_2 \ln \delta - \lambda(1 + \delta) \cdot 30$$

$$\frac{\partial l}{\partial \lambda} = \frac{N_1 + N_2}{\lambda} - (1 + \delta) \cdot 30$$

$$\frac{\partial l}{\partial \delta} = \frac{N_2}{\delta} - \lambda \cdot 30$$

Solve likelihood equations $\frac{\partial l}{\partial \lambda} = \frac{\partial l}{\partial \delta} = 0$

gives $\hat{\lambda}(1) = \frac{N_1}{30} = \frac{4}{30} = 0.1333$

$$\hat{\delta}(1) = \frac{N_2}{N_1} = \frac{7}{4} = 1.75$$

and $l(\hat{\lambda}(1), 1, \hat{\delta}(1)) = 11 \cdot \ln\left(\frac{4}{30}\right) + 7 \cdot \ln\left(\frac{7}{4}\right) - \frac{4}{30} \left(1 + \frac{7}{4}\right) 30$
 $= -29.24662$

So $W = 2 [-26.88 - (-29.25)] = 4.73$

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Reject at 5% if $W \geq 3.84$, so we reject!