

TMA4275 LIFETIME ANALYSIS

Slides 4: Gumbel distribution. Log-location-scale families

Bo Lindqvist
Department of Mathematical Sciences
Norwegian University of Science and Technology
Trondheim

<http://www.math.ntnu.no/~bo/>
bo@math.ntnu.no

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EXTREME VALUE DISTRIBUTIONS

Let T_1, T_2, \dots, T_n be lifetimes of n components, with ordered values denoted by $T_{(1)} < T_{(2)} < \dots < T_{(n)}$. Thus $T_{(1)}$ is the minimum and corresponds to the lifetime of a series system.

For large n , $T_{(1)}$ is approximately Weibull-distributed. *This motivates the widespread use of the Weibull-distribution!*

If the T_i are no longer lifetimes, but have support in $(-\infty, \infty)$, then the limiting distribution of a properly normalized version of $T_{(1)}$ equals the distribution of a random variable Y with cdf

$$F_Y(y) = 1 - e^{-e^{\frac{y-\mu}{\sigma}}}, \quad -\infty < y < \infty$$

This is the so called “Distribution of smallest extreme”, or “Extreme value distribution of type I”, or (which we will call it) the **Gumbel-distribution**; $Y \sim \text{Gumbel}(\mu, \sigma)$.

We write $Y \sim \text{Gumbel}(\mu, \sigma)$

Show that

$$F_Y(y) = 1 - e^{-e^{\frac{y-\mu}{\sigma}}}, \quad -\infty < y < \infty$$

satisfies the requirements for a cdf, i.e.

- Increasing in y
- $\lim_{y \rightarrow -\infty} F_Y(y) = 0$
- $\lim_{y \rightarrow \infty} F_Y(y) = 1$

WHY ARE WE INTERESTED IN THE GUMBEL DISTRIBUTION?

If T is Weibull-distributed, $T \sim \text{Weib}(\alpha, \theta)$, then $Y = \ln T$ is Gumbel-distributed, $Y \sim \text{Gumbel}(\mu, \sigma)$, with $\mu = \ln \theta$, $\sigma = 1/\alpha$.

Proof: Note first that $T = e^Y$ and $R(t) = P(T > t) = e^{-\left(\frac{t}{\theta}\right)^\alpha}$. Then:

$$\begin{aligned}P(Y > y) &= P(e^Y > e^y) = P(T > e^y) = R(e^y) \\&= e^{-\left(\frac{e^y}{\theta}\right)^\alpha} = e^{-\left(\frac{e^y}{e^{\ln \theta}}\right)^\alpha} \\&= e^{-(e^{y-\ln \theta})^\alpha} = e^{-e^{\left(\frac{y-\ln \theta}{1/\alpha}\right)}}\end{aligned}$$

Thus, $F_Y(y) = 1 - P(Y > y) = 1 - e^{-e^{\left(\frac{y-\ln \theta}{1/\alpha}\right)}}$, which shows that $Y \sim \text{Gumbel}(\ln \theta, 1/\alpha)$.

We shall see later why this is a useful and interesting result (and not just a curiosity...)

Let $Y \sim \text{Gumbel}(\mu, \sigma)$ and recall the cdf

$$F_Y(y) = P(Y \leq y) = 1 - e^{-e^{\frac{y-\mu}{\sigma}}} \quad \text{for } -\infty < y < \infty$$

The cdf of $\text{Gumbel}(0,1)$, called *the standard Gumbel distribution*, is

$$G(w) = 1 - e^{-e^w} \quad \text{for } -\infty < w < \infty$$

It is seen that

$$F_Y(y) = P(Y \leq y) = G\left(\frac{y - \mu}{\sigma}\right)$$

This defines the cdf of the $\text{Gumbel}(\mu, \sigma)$ in terms of the cdf of the standard Gumbel, in the same way as the cdf of $Y \sim N(\mu, \sigma)$ can be expressed by the cdf, $\Phi(\cdot)$, of the standard normal. How?

Let again $Y \sim \text{Gumbel}(\mu, \sigma)$, and define

$$W = \frac{Y - \mu}{\sigma} \quad (*)$$

Then W has the standard Gumbel distribution. This is seen as follows:

$$\begin{aligned} P(W \leq w) &= P\left(\frac{Y - \mu}{\sigma} \leq w\right) \\ &= P(Y \leq \mu + \sigma w) = F_Y(\mu + \sigma w) \\ &= 1 - e^{-e^w} \equiv G(w) \end{aligned}$$

By solving () for Y it follows that we have the representation for $Y \sim \text{Gumbel}(\mu, \sigma)$:*

$$Y = \mu + \sigma W$$

where $W \sim \text{Gumbel}(0, 1)$.

Recall once more that if $W \sim \text{Gumbel}(0, 1)$, then W has the cdf

$$G(w) = 1 - e^{-e^w}$$

The pdf of W is hence

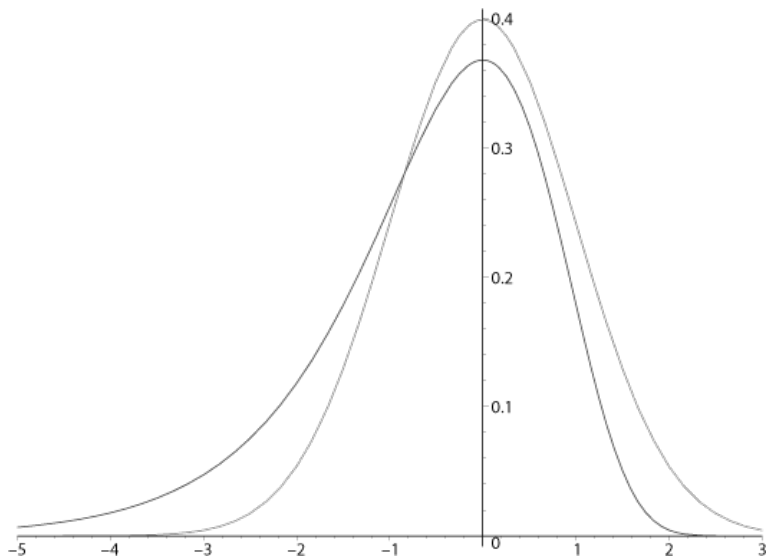
$$g(w) = G'(w) = -e^{-e^w} (-e^w) = e^w e^{-e^w}$$

We also have

$$E(W) = \int_{-\infty}^{\infty} w e^w e^{-e^w} dw = -\gamma,$$

where $\gamma = -0.5772$ is *Euler's constant*.

STANDARD GUMBEL AND NORMAL DISTRIBUTIONS



We have seen:

$$T \sim \text{lognorm}(\mu, \sigma) \iff Y = \ln T \sim N(\mu, \sigma)$$

$$T \sim \text{Weib}(\alpha, \theta) \iff Y = \ln T \sim \text{Gumbel}(\mu, \sigma), \text{ with } \mu = \ln \theta, \sigma = 1/\alpha.$$

- Both distributions thus define **log-location-scale families**, which are characterized by the fact that $Y = \ln T$ has a cdf which can be expressed as

$$F_Y(y) = P(Y \leq y) = \Psi\left(\frac{y - \mu}{\sigma}\right)$$

where $\Psi(\cdot)$ is the cdf of some “standardized distribution” on $(-\infty, \infty)$.

- Equivalently, log-location-scale families are characterized by representations

$$\ln T = \mu + \sigma U$$

where U has cdf $\Psi(\cdot)$ as described above.

Generally, $\mu \in (-\infty, +\infty)$ is called the *location parameter*, and $\sigma > 0$ is called the *scale parameter*.

A random variable Y has the **logistic distribution** with location parameter μ and scale parameter σ , written $Y \sim \text{logistic}(\mu, \sigma)$, if

$$F_Y(y) = P(Y \leq y) = H\left(\frac{y - \mu}{\sigma}\right) \quad \text{for } -\infty < y < \infty$$

where

$$H(v) = P(V \leq v) = \frac{e^v}{1 + e^v} \quad \text{for } -\infty < v < \infty$$

is the cdf of the standard logistic distribution, $\text{logistic}(0, 1)$.

A lifetime T has the **log-logistic** distribution with location parameter μ and scale parameter σ if $Y = \ln T \sim \text{logistic}(\mu, \sigma)$. In this case we have the representation

$$\ln T = \mu + \sigma V$$

where $V \sim \text{logistic}(0, 1)$.

Recall that if $V \sim \text{logistic}(0, 1)$, then the cdf of V is $H(v) = P(V \leq v) = \frac{e^v}{1+e^v}$ for $-\infty < v < \infty$.

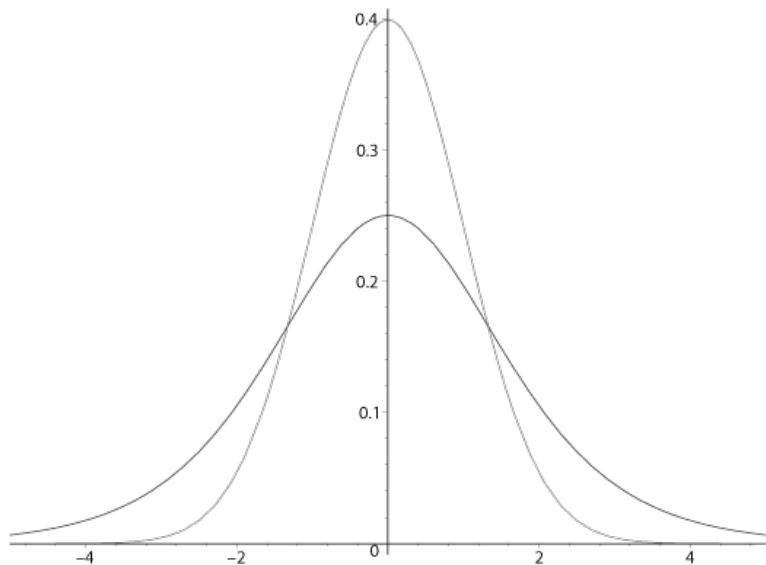
Hence the pdf of V is

$$h(v) = H'(v) = \frac{e^v}{(1 + e^v)^2} \quad (\text{do the differentiation!})$$

Like the standard normal, this density is symmetric around the y -axis (which is not the case for the standard Gumbel).

Check this by showing that $h(-v) = h(v)$ for all v .

STANDARD LOGISTIC AND STANDARD NORMAL



By assumption, $Y = \ln T$ has a cdf which can be expressed as

$$F_Y(y) = P(Y \leq y) = \Psi \left(\frac{y - \mu}{\sigma} \right) \quad \text{for } -\infty < y < \infty$$

where $\Psi(\cdot)$ is the cdf of a standard distribution. Let further $\psi(u) = \Psi'(u)$.

Then

$$R(t) = P(T > t) = P(\ln T > \ln t) = 1 - \Psi \left(\frac{\ln t - \mu}{\sigma} \right)$$

$$f(t) = -R'(t) = \psi \left(\frac{\ln t - \mu}{\sigma} \right) \cdot \frac{1}{t\sigma}$$

$$z(t) = \frac{f(t)}{R(t)} = \frac{\psi \left(\frac{\ln t - \mu}{\sigma} \right) / (t\sigma)}{1 - \Psi \left(\frac{\ln t - \mu}{\sigma} \right)}$$

(as already obtained for the lognormal distribution).