TMA4275 LIFETIME ANALYSIS

Slides 16-DRAFT: Parametric estimation in NHPPs

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> > NTNU, Spring 2014

Simple Example With 3 Systems

Power Law NHPP Model: $W(t; \alpha, \theta) = (t/\theta)^{\alpha}$

Results for: SimpleNHPP.MTW

Parametric Growth Curve: Time

System: ID

Model: Power-Law Process

Estimation Method: Maximum Likelihood

Parameter Estimates

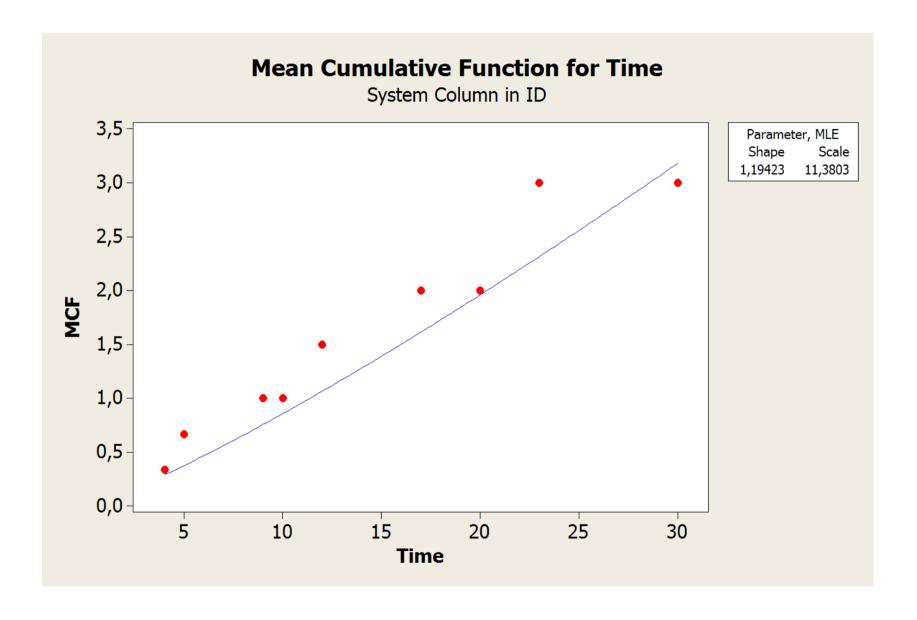
		Standard	95% Nor	mal CI
Parameter	Estimate	Error	Lower	Upper
Shape	1,19423	0,445	0,323015	2,06545
Scale	11,3803	4,840	1,89335	20,8672

Test for Equal Shape Parameters
Bartlett's Modified Likelihood Ratio Chi-Square

Test Statistic 0,06 P-Value 0,972 DF 2

	MIL-Hdbk-189		Laplac	e's		
	TTT-based	Pooled	TTT-based	Pooled	Anderson-Darling	
Test Statistic	9,03	8,89	0,28	0,31	0,28	
P-Value	0,599	0,576	0,781	0,756	0,954	
DF	12	12				

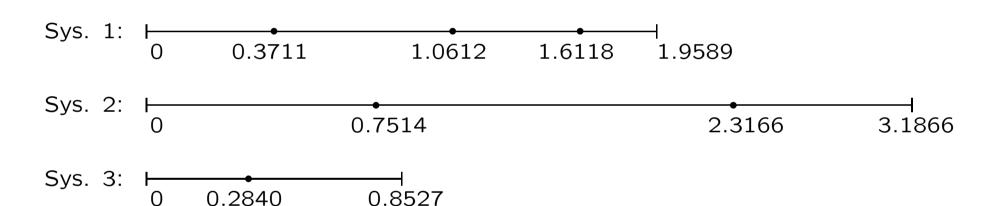
Simple Example With 3 Systems



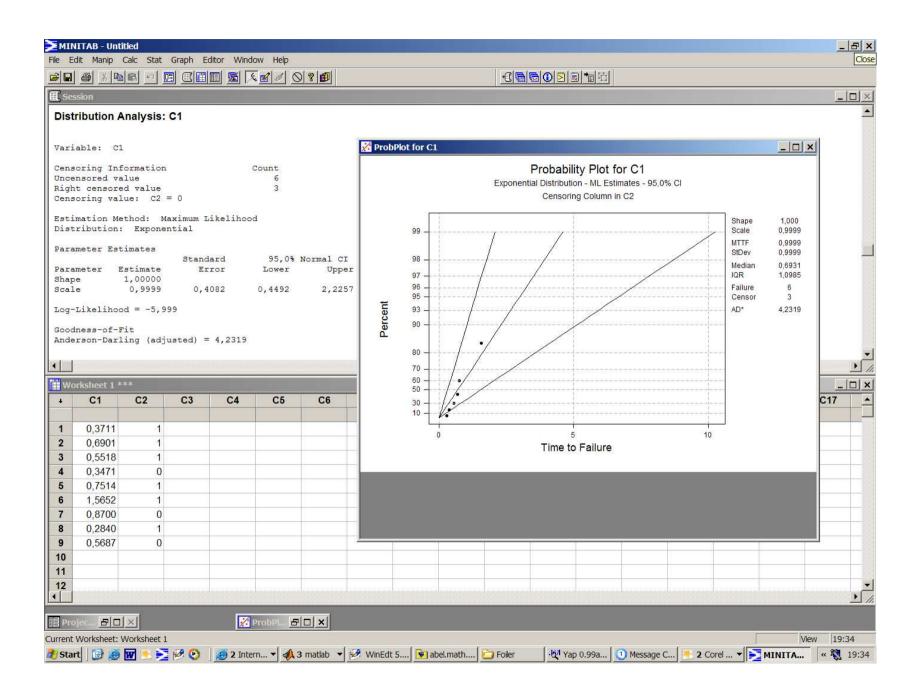
RESIDUAL PROCESS: "SIMPLE EXAMPLE".

Data points (and endpoints on axes) are transformed with the estimated cumulative ROCOF,

$$\hat{W}(t) = 0.0538 \cdot t^{1.20}$$



Times between events, plus censored times at the end of each axis, are on the next slide anlysed by MINITAB as a set of censored exponential variables.

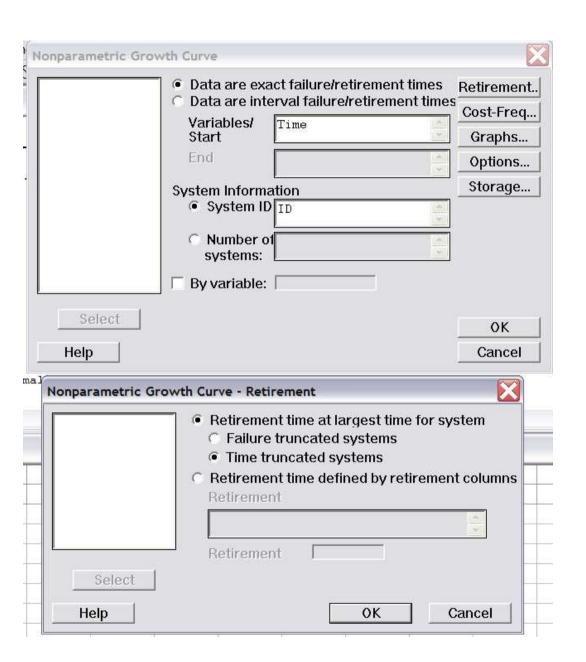


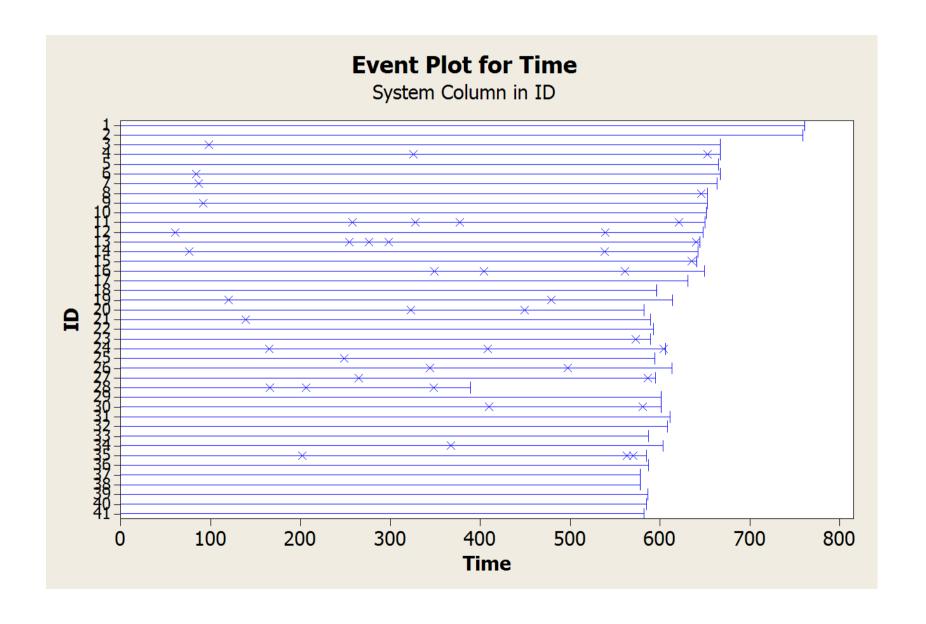
Valve Seat Replacement Times (Nelson and Doganaksoy 1989)

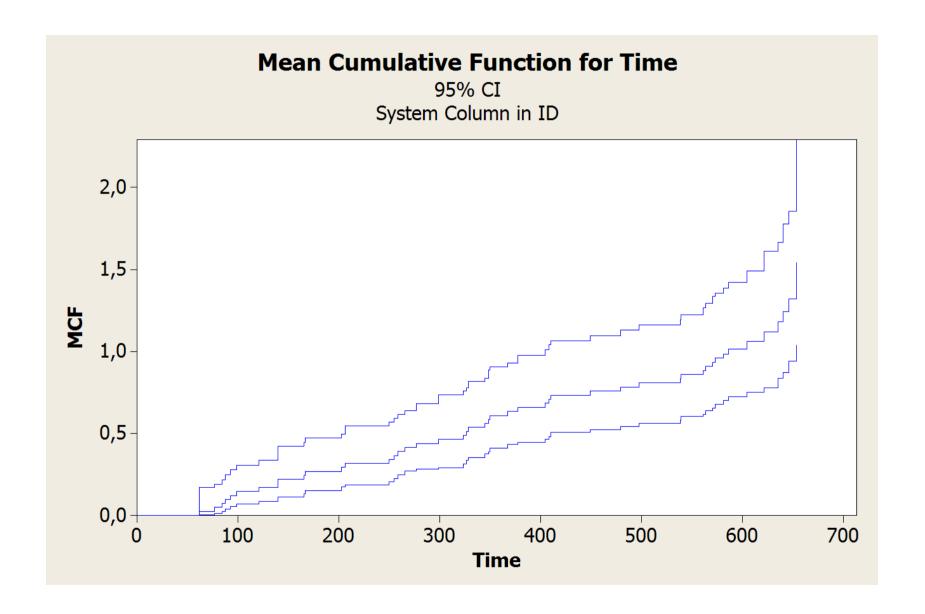
Data collected from valve seats from a fleet of 41 diesel engines (days of operation)

- Each engine has 16 valves
- Does the replacement rate increase with age?
- How many replacement valves will be needed in the future?
- Can valve life in these systems be modeled as a renewal process?

+	C1	C2	C3	C4	C5	C6	C7	C8	
	ID	Time							
1	1	761							
2	2	759							
3	3	98							
4	3	667							
5	4	326							
6	4	653							
7	4	653							
8	4	667							
9	5	665							
10	6	84							
11	6	667							
12	7	87							
13	7	663							
14	8	646							
15	8	653							
16	9	92							
17	9	653							
18	10	651							
19	11	258							
20	11	328							
21	11	377							
22	11	621							
23	11	650							
24	12	61							
25	12	539							
26	12	648							







Nonparametric Growth Curve: Time

System: ID

Nonparametric Estimates

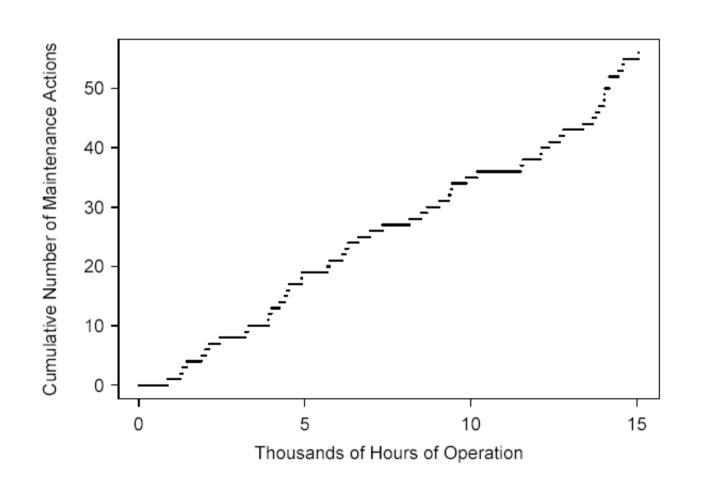
Table of Mean Cumulative Function

	Mean				
	Cumulative	Standard	95% No	rmal CI	
Time	Function	Error	Lower	Upper	System
61	0,02439	0,024091	0,00352	0,16903	12
76	0,04878	0,033641	0,01262	0,18848	14
84	0,07317	0,040670	0,02462	0,21750	6
87	0,09756	0,046340	0,03846	0,24750	7
92	0,12195	0,051105	0,05364	0,27726	9
98	0,14634	0,055199	0,06987	0,30650	3
120	0,17073	0,058764	0,08696	0,33519	19
139	0,19512	0,061891	0,10479	0,36333	21
139	0,21951	0,073270	0,11411	0,42226	21
165	0,24390	0,075417	0,13305	0,44711	24
166	0,26829	0,077317	0,15251	0,47196	28
202	0,29268	0,078988	0,17246	0,49672	35
206	0,31707	0,087527	0,18458	0,54467	28
249	0,34146	0,088680	0,20525	0,56807	25
254	0,36585	0,089656	0,22631	0,59143	13
258	0,39024	0,090461	0,24775	0,61468	11
265	0,41463	0,091101	0,26955	0,63780	27
276	0,43902	0,097858	0,28363	0,67955	13
298	0,46341	0,109607	0,29150	0,73671	13
323	0,48780	0,109740	0,31387	0,75812	20
326	0,51220	0,109740	0,33656	0,77949	4
328	0,53659	0,114907	0,35266	0,81643	11
344	0,56098	0,114654	0,37581	0,83737	26
348	0,58537	0,124250	0,38615	0,88737	28
349	0,60976	0,123782	0,40960	0,90772	16
367	0,63415	0,123194	0,43334	0,92801	34
377	0,65854	0,131842	0,44480	0,97498	11
404	0,68354	0,135939	0,46289	1,00936	16

Times of Unscheduled Maintenance Actions for a USS Grampus Diesel Engine

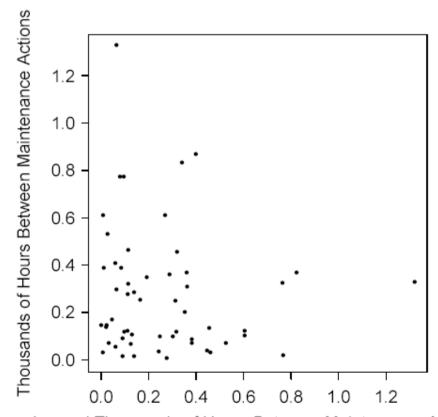
- Unscheduled maintenance actions caused by failure of imminent failure.
- Unscheduled maintenance actions are inconvenient and expensive.
- Data available for 16,000 operating hours.
- Data from Lee (1980).
- Is the system deteriorating (i.e., are failures occurring more rapidly as the system ages)?
- Can the occurrence of unscheduled maintenance actions be modeled by an HPP?

Cumulative Number of Unscheduled Maintenance Actions Versus Operating Hours for a USS Grampus Diesel Engine Lee (1980)



Grampus- data: Plot of (T_i, T_{i+1}) to investigate whether times between failures can be assumed independent. The figure does not indicate a correlation between successive times.

USS Grampus Diesel Engine Plot of Times Between Unscheduled Maintenance Actions Versus Lagged Times Between Unscheduled Maintenance Actions



The Likelihood for the NHPP - Single Unit

With interval recurrence data.

Suppose that the unit has been observed for a period $(0, t_a]$ and the data are the number of recurrences d_1, \ldots, d_m in the nonoverlapping intervals $(t_0, t_1], (t_1, t_2], \ldots, (t_{m-1}, t_m]$ (with $t_0 = 0$, $t_m = t_a$).

$$L(\theta) = \Pr[N(t_0, t_1) = d_1, \dots, N(t_{m-1}, t_m) = d_m]$$

$$= \prod_{j=1}^{m} \Pr[N(t_{j-1}, t_j; \theta)]^{d_j}$$

$$= \prod_{j=1}^{m} \frac{\left[\mu(t_{j-1}, t_j; \theta)\right]^{d_j}}{d_j!} \exp\left[-\mu(t_{j-1}, t_j; \theta)\right]$$

$$= \prod_{j=1}^{m} \frac{\left[\mu(t_{j-1}, t_j; \theta)\right]^{d_j}}{d_j!} \times \exp\left[-\mu(t_0, t_a; \theta)\right]$$

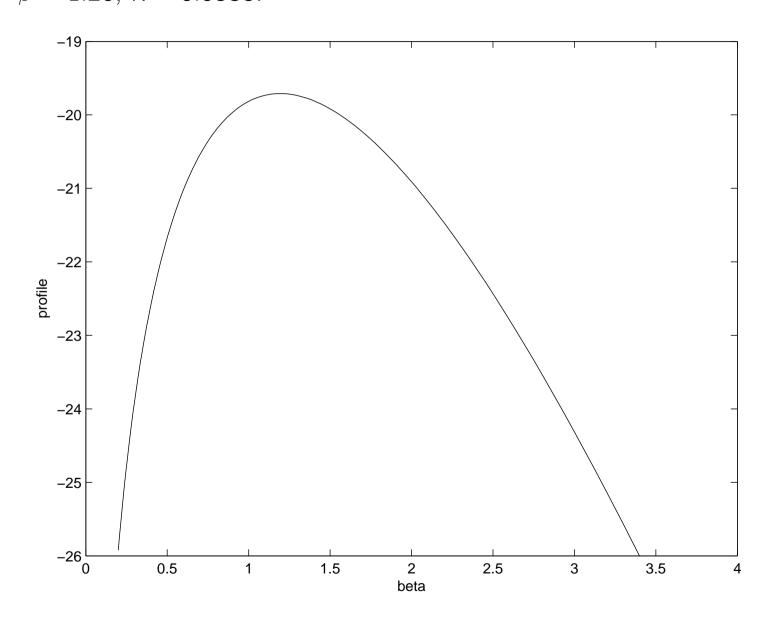
The Likelihood for the NHPP (Continued)

• If the number of intervals m increases and there are **exact** recurrences at $t_1 \leq \ldots \leq t_r$ (here $r = \sum_{j=1}^m d_j$, $t_0 \leq t_1$, $t_r \leq t_a$), then using a limiting argument it follows that the likelihood in terms of the density approximation is

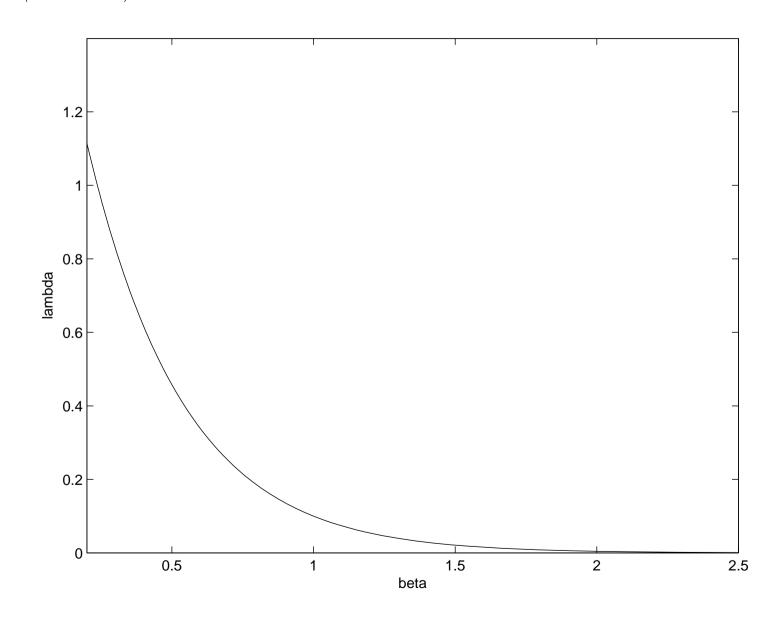
$$L(\boldsymbol{\theta}) = \prod_{j=1}^{r} \nu(t_j; \boldsymbol{\theta}) \times \exp\left[-\mu(0, t_a; \boldsymbol{\theta})\right]$$

- For simplicity, above we assumed that the intervals are contiguous. Obvious changes to the formula above give the likelihood when there are gaps among the intervals.
- In both cases (the interval data or exact recurrences data) the same methods used in Chapters 7, 8 can be used to obtain the ML estimate $\hat{\theta}$ and confidence regions for θ or functions of θ .

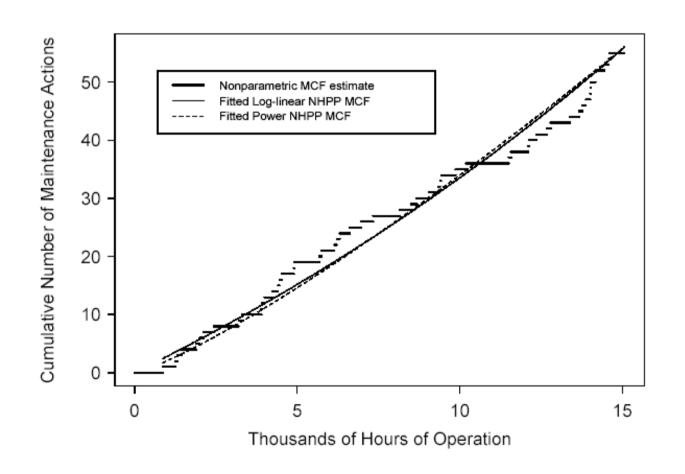
PROFILE LIKELIHOOD FOR BETA ("SIMPLE EXAMPLE") $\hat{\beta}=1.20,\;\hat{\lambda}=0.0538.$



CONNECTION BETWEEN LAMBDA OG BETA ("SIMPLE EXAMPLE") $\hat{\beta}=1.20,\;\hat{\lambda}=0.0538.$



Cumulative Number of Unscheduled Maintenance Actions Versus Operating Hours with Power and Loglinear NHPP Models for a USS Grampus Diesel Engine



Results of Fitting NHPP Models to the USS Grampus Diesel Engine Data

- Both models seem to fit the data very well.
- For the power recurrence rate model, $\hat{\beta}$ =1.22 and $\hat{\eta}$ =0.553.
- For the loglinear recurrence rate model, $\hat{\gamma}_0$ =1.01 and $\hat{\gamma}_1$ =.0377.
- Times between recurrences are consistent with a HPP:
 - ▶ the Lewis-Robinson test gave $Z_{LR} = 1.02$ with p-value p = .21.
 - ▶ the MIL-HDBk-189 test gave $X_{\text{MHB}}^2 = 92$ with p-value p = .08.

Comparison of trend tests

Minitab provides five trend tests for data with multiple systems: MIL-hdbk-189 (TTT-based), MIL-hdbk-189 (Pooled), Laplace's (TTT-based), Laplace's (Pooled), and Anderson-Darling. The pooled Laplace and military handbook tests reduce to their respective TTT-based tests when there is only one system. These tests behave differently under the following two circumstances:

- 1 the data follow a non-monotonic trend
- 2 the data are from heterogeneous systems

Monotonic and non-monotonic trends

There is a trend in the pattern of times between failure if the times change in a systematic way. Trends can be:

- monotonic times between failures are getting either consistently longer (decreasing trend) or consistently shorter (increasing trend)
- non-monotonic times between failures alternate between increasing and decreasing trend (cyclic) or have a decreasing trend, no trend, and then increasing trend (bathtub)

The Anderson-Darling test will reject the null hypothesis in the presence of both monotonic and non-monotonic trends. The other tests will generally only detect monotonic trends. While the Anderson-Darling test is useful if you suspect the existence of a cyclic or other non-monotonic trend, the other tests are more powerful in the case of a monotonic trend.

Homogeneous and heterogeneous systems

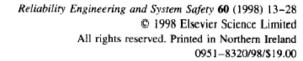
The null hypothesis of no trend differs slightly for the different tests:

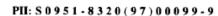
- The null hypothesis for the pooled tests (MIL-hdbk-189 and Laplace's) is that the data come from a homogeneous Poisson processes (HPP) with a possibly different MTBF for each system. Thus, rejecting the null hypothesis means that you can definitely conclude there is a trend in your data.
- The null hypothesis for the TTT-based tests (MIL-hdbk-189, Laplace's, and Anderson-Darling) is that the data come from a homogeneous Posson process (HPP) with the same MTBF for each system. Thus, rejecting the null hypothesis could mean that either there is a trend in your data or your data come from heterogeneous systems. Therefore, you should use TTT-based tests only when you are confident that your systems are homogeneous.

The table below summarizes the different null hypotheses associated with the trend tests.

	MIL-hdbk-189 (Pooled)	MIL-hdbk-189 (TTT-based)	Laplace's (Pooled)	Laplace's (TTT-based)	Anderson- Darling
Null Hypothesis	HPP (possibly different MTBFs)	HPP (equal MTBFs)	HPP (possibly different MTBFs	HPP (equal)MTBFs)	HPP (possibly different MTBFs)
Rejecting H ₀ means	monotonic trend	monotonic trend or systems are heterogeneous	monotonic trend	monotonic trend or systems are heterogeneous	

See 12 for more information concerning these tests.





TTT-based tests for trend in repairable systems data

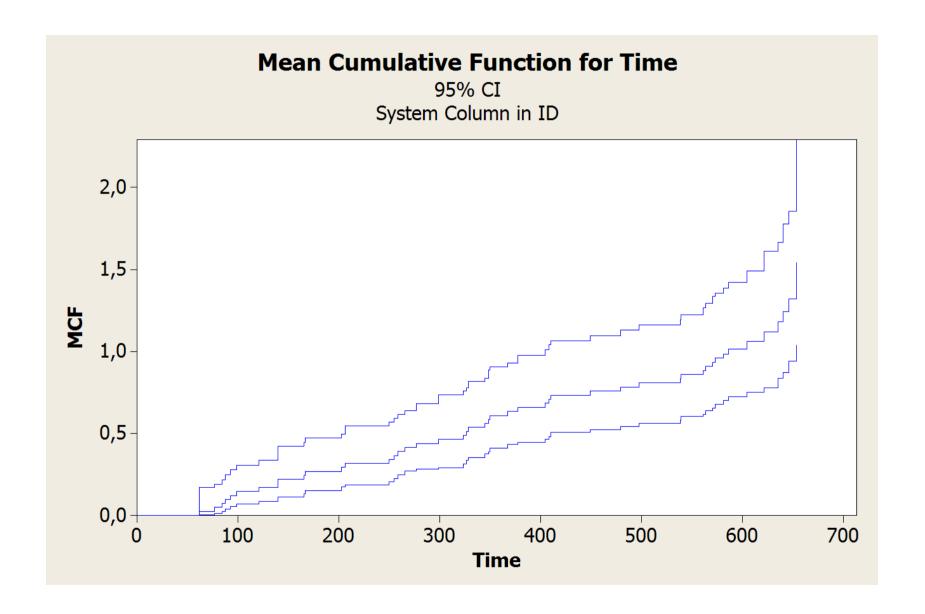
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(Received 25 September 1996; revised 24 January 1997; accepted 15 July 1997)

A major aspect of analysis of failure data for repairable systems is the testing for a possible trend in interfailure times. This paper reviews some important and popular graphical methods and tests for the nonhomogeneous Poisson process model. In particular, the total time on test (TTT) plot is considered, and trend tests based on the TTT-statistic are motivated and derived. In particular, a test based on the Anderson–Darling statistic is suggested. The tests are evaluated and compared in a simulation study, both with respect to the achievement of correct significance level and rejection power. The considered alternatives to 'no trend' are the log-linear, power law and a class of bathtub-shaped intensity functions. The simulation study involves single systems, as well as the case where several independent systems of the same kind are observed. © 1998 Elsevier Science Limited.

Valveseat Data



Valveseat Data

	MIL-Hdbk-189		Laplace's			
	TTT-based	Pooled	TTT-based	Pooled	Anderson-Darling	
Test Statistic	80,28	66,15	0,46	2,38	0,80	
P-Value	0,249	0,017	0,645	0,017	0,478	
DF	96	96				

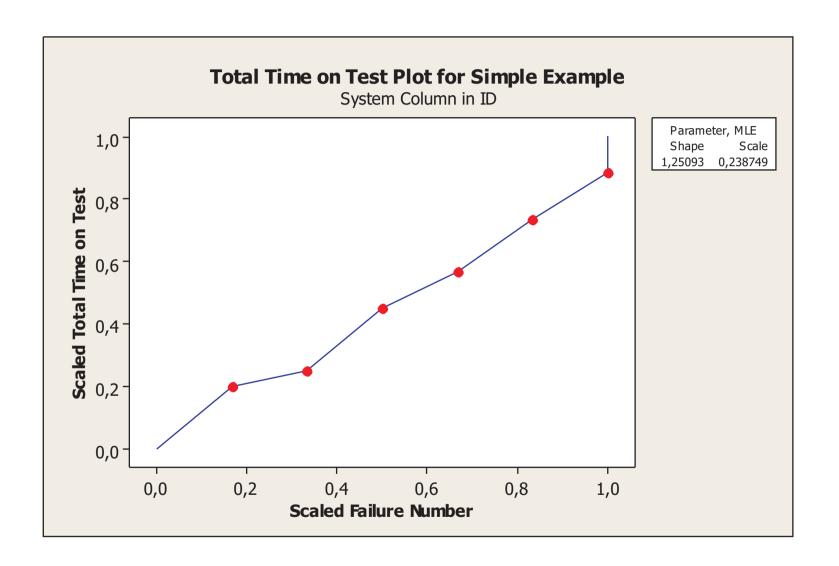
TTT-analysis Simple Example

Row	STTT	${ t ID}$	Scaled
1	12	1	0,20000
2	15	1	0,25000
3	27	1	0,45000
4	34	1	0,56667
5	44	1	0,73333
6	53	1	0,88333
7	60	1	1,00000

Parameter Estimates

		Standard	95% Nor	mal CI
Parameter	Estimate	Error	Lower	Upper
Shape	1,25093	0,511	0,249996	2,25186
Scale	0,238749	0,160	-0,0746105	0,552109

	MIL-Hdbk-189	Laplace's	Anderson-Darling
Test Statistic	9,59	0,12	0,24
P-Value	0,697	0,906	0,977
DF	12		



TTT-analysis of Valve Seat Data

Parametric Growth Curve: C1

Model: Power-Law Process

Estimation Method: Maximum Likelihood

Parameter Estimates

		Standard	95% Nor	mal CI
Parameter	Estimate	Error	Lower	Upper
Shape	1,39706	0,202	1,00184	1,79229
Scale	0,0626023	0,026	0,0119179	0,113287

	MIL-Hdbk-189	Laplace's	Anderson-Darling
Test Statistic	68,72	2,03	3,17
P-Value	0,032	0,043	0,022
DF	96		

