

# STK4080 SURVIVAL AND EVENT HISTORY ANALYSIS

## Slides 3: Censoring

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Lifetime data typically include *censored* data, meaning that:

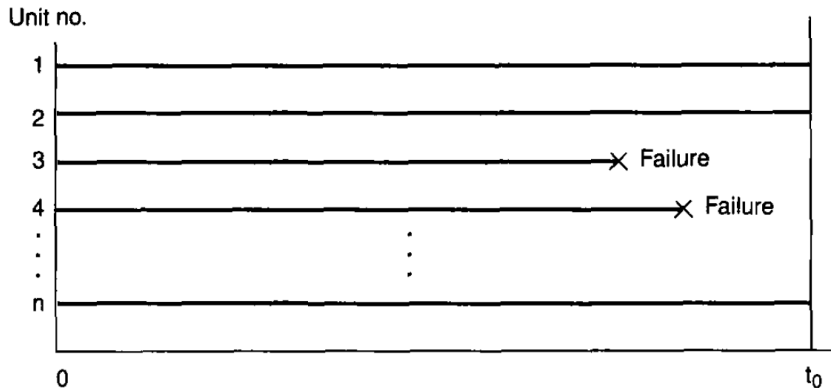
- some lifetimes are known to have occurred only within certain intervals.
- The remaining lifetimes are known exactly.

*Categories of censoring:*

- right censoring
- left censoring
- interval censoring

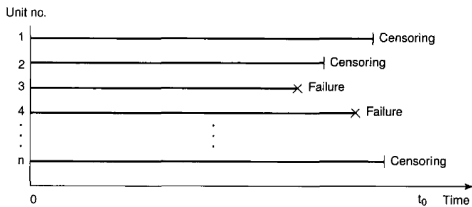
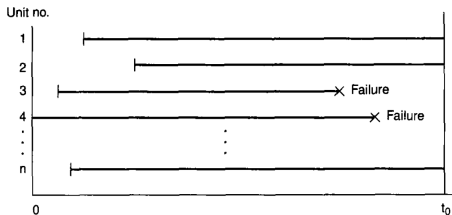
# TYPE I (RIGHT) CENSORING

$n$  units put on test at time  $t = 0$ . Experiment stopped at time  $t = t_0$ .



# GENERALIZED TYPE I CENSORING ("STAGGERED ENTRY")

Individuals enter the study at different times, and the terminal point of the study is predetermined.



$n$  units are put on test at time  $t = 0$ .

The study continues until  $r$  individuals have failed, where  $r$  is some predetermined integer ( $r < n$ ).

*Advantage:* It could take a very long time for all items to fail. Also, the statistical treatment of Type II censored data is simpler because the joint distribution of the order statistics is available.

This is a mix of Type I and Type II censoring. Choose both an end time  $t_0$  as for Type I censoring and an  $r < n$  as for Type II censoring. Stop the experiment at time  $t_0$  or at the  $r$ th failure, whatever comes first.

- For each unit we define
  - $T_i$  to be the potential lifetime
  - $C_i$  to be the potential censoring time

where

- $T_i, C_i$  are **independent random variables**.
- Then we *observe* the pair  $(\tilde{T}_i, D_i)$ , where

$$\begin{aligned}\tilde{T}_i &= \min(T_i, C_i) \\ D_i &= \begin{cases} 1 & \text{if } T_i \leq C_i \\ 0 & \text{if } T_i > C_i \end{cases}\end{aligned}$$

*Example of use:* Cancer treatment, with  $T_i$  being the time of death due to this cancer; while  $C_i$  is the time of death of another cause, or an accident, or migration, etc.

(*Right censoring is the most common way of censoring.* )

Right censoring of Type I, II, III, IV can all be represented as follows:

$n$  units are observed, with potential i.i.d. lifetimes  $T_1, T_2, \dots, T_n$ . For each  $i$ , we observe a time  $\tilde{T}_i$  which is either the true lifetime  $T_i$ , or a censoring time  $C_i < T_i$ , in which case the true lifetime is “to the right” of the observed time  $C_i$ .

The observation from a unit is the pair  $(\tilde{T}_i, D_i)$  where the *censoring indicator*  $D_i$  is defined by

$$D_i = \begin{cases} 1 & \text{if } \tilde{T}_i = T_i, \text{ in which case we observe the true lifetime } T_i \\ 0 & \text{if } \tilde{T}_i = C_i, \text{ in which case it is only known that } T_i > Y_i \end{cases}$$



Consider a situation where  $n$  individuals are followed from time  $t = 0$ . The  $i$ th individual is followed until  $\tilde{T}_i = \min(T_i, C_i)$ , i.e. until either failure (death) or censoring at time  $C_i$ .

*The  $i$ th individual is said to be at risk at time  $t$  if  $t < \tilde{T}_i$ , i.e. if the individual has not yet been censored and have not failed.*

A censoring scheme is said to satisfy the property of **independent censoring** if, at any time  $t$ , the individuals that are *at risk* are representative for the distribution of  $T$  in the sense that their probability of failing in a small time interval  $(t, t + h)$  is (in the limit as  $h$  tends to 0) is  $\alpha(t)h$ .

*The censoring types we have considered so far all satisfy this independent censoring property.*