Hazard rate, cumulative hazard and survival function in the absolutely continuous case

Let T be a lifetime with an absolutely continuous distribution. Then

$$\alpha(t) = \frac{f(t)}{S(t)} = -\frac{S'(t)}{S(t)}$$

$$A(t) = \int_0^t \alpha(s)ds$$

$$S(t) = e^{-A(t)}$$

Hazard rate and cumulative hazard in the general case

A general definition of hazard and cumulative hazard were given in Slides 8, starting from a general S(t) = P(T > t) only assuming

- \triangleright S(t) is continuous from the right
- \triangleright S(t) is non-increasing
- 0 < S(0) < 1

Then it was shown that

$$dA(t) = P(t \le T < t + dt | T \ge t) = -\frac{dS(t)}{S(t-)}$$

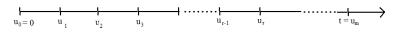
leading to the general expression

$$A(t) = -\int_0^t \frac{dS(u)}{S(u-)}$$

As a check, for the absolutely continuous case we would then have

$$A(t) = -\int_0^t \frac{dS(u)}{S(u-)} = \int_0^t \frac{-S'(u)du}{S(u)} = \int_0^t \frac{f(u)du}{S(u)} = \int_0^t \alpha(u)du$$

The discrete hazard and cumulative hazard from scratch



Let $u_0 = 0, u_1, u_2, \ldots$ be the values of T.

Reconsider $dA(t) = P(t \le T \le t + dt | T \ge t)$. Looking at the above axis it is clear that this is 0 unless t is one of the u_j .

For $t = u_i$, we get on the other hand,

$$dA(t) = P(t \le T < t + dt | T \ge t) = P(T = u_j | T \ge u_j) \equiv \alpha_{u_j}$$

which we called the discrete hazard rate.

Thus $A(t) = \int_0^t dA(u)$ is a function which jumps by α_{u_j} when $t = u_j$, so we get the natural result that

$$A(t) = \sum_{u_i < t} \alpha_{u_j}$$

The discrete survival function and the product-integral

For $u_m \le t < u_{m+1}$ we now have

$$S(t) = P(T > t) = P(T > u_m)$$

$$= \prod_{j=1}^{m} P(T > u_j | T > u_{j-1}) = \prod_{j=1}^{m} (1 - P(T = u_j | T > u_{j-1}))$$

$$= \prod_{i=1}^{m} (1 - \alpha_{u_i}) = \prod_{i=1}^{m} (1 - dA(u_i)) = \prod_{u_i < t} (1 - dA(u_i)) \quad (*)$$

This is a special case of the so called product-integral

$$S(t) = \prod_{0 \le u \le t} (1 - dA(u))$$

When A(t) makes jumps at discrete time points u_j , this is exactly given as in (*).

The general product-integral

Recall the general expressions

$$dA(u) = -\frac{dS(u)}{S(u-)}, \quad A(t) = -\int_0^t \frac{dS(u)}{S(u-)}$$

Solving S(t) from these equations in the general case, leads to the so-called ${\bf product\text{-}integral}$,

$$S(t) = \prod_{0 \le u \le t} (1 - dA(u))$$

which can be defined as a limit as follows:

Let t>0 be fixed, and let $0=u_0< u_1< u_2<\ldots< u_m\equiv t$ define a partition of [0,t]. Let $m\to\infty$ in a way such that the spacings u_j-u_{j-1} tend to 0. Then

$$S(t) = \prod_{0 \le u \le t} (1 - dA(u)) =_{def} \lim_{m \to \infty} \prod_{j=1}^{m} [1 - (A(u_j) - A(u_{j-1}))]$$

A quick check for the continuous case

Suppose
$$A(t) = \int_0^t \alpha(u) du$$
 so that $A'(t) = \alpha(t)$.

Then for a given partition with large m and fixed $t = u_m$,

$$\prod_{j=1}^{m} [1 - (A(u_j) - A(u_{j-1}))] \approx \prod_{j=1}^{m} (1 - \alpha(u_j)(u_j - u_{j-1}))$$

$$\approx \prod_{j=1}^{m} e^{-\alpha(u_j)(u_j - u_{j-1})} = e^{-\sum_{j=1}^{m} \alpha(u_j)(u_j - u_{j-1})}$$

$$\approx e^{-\int_0^t \alpha(u) du} \equiv S(t)$$

Thus, the product-integral gives the correct result for continuous distribution, as we saw that it did for the discrete case.

The Kaplan-Meier estimator

Bo Lindqvist Slides 8: Kaplan-Meier

Let's go back to the discrete survival function,

$$S(t) = \prod_{u_j \leq t} (1 - \alpha_{u_j}) = \prod_{u_j \leq t} (1 - dA(u_j))$$

An estimator of S(t) can hence be obtained by putting an estimate for the function A(t) in the above. This may be done by the Nelson-Aalen estimator and turns out to lead to the Kaplan-Meier estimator.

Recall the NA-estimator $\hat{A}(t) = \int_0^t \frac{dN(s)}{Y(s)}$. This is a discrete function with jumps $\frac{dN(s)}{Y(s)}$, i.e., it jumps at the failure times T_i . Thus we can write

$$\hat{S}(t) = \prod_{0 \le u \le t} (1 - d\hat{A}(u)) = \prod_{0 \le u \le t} \left(1 - \frac{dN(u)}{Y(u)} \right) = \prod_{T_i \le t} \left(1 - \frac{1}{Y(T_i)} \right)$$