#### SOLUTION OBLIGATORY STK4080/9080 SPRING 2021

First we read the data into R by:

popes = read.csv("https://folk.ntnu.no/bo/STK4080/2021/popes\_25\_December\_2016.csv")

In order to have updated data, knowing that both Benedict XVI and Francis are still alive, we add the survival time from 25.12.2016 to 01.03.2021 for these popes, i.e., 4.18 years. This gives us the new data frame popes.new, obtained as follows:

```
popes.new = popes
popes.new$Survival[1]=3.79+4.18
popes.new$Survival[2]=11.70+4.18
```
As a check, you get the information of the last six popes by the command

```
head(popes.new)
```
We then introduce the working variables:

time = popes.new\$Survival delta = 1 - popes.new\$Censored ageelect = popes.new\$Age.Election yearelect = popes.new\$Year.Elected

#### **PROBLEM 1**

#### **(a)**

Number of popes: 63

> summary(time)





delta =  $1 : 61$  popes delta =  $0 : 2$  popes

Plots:

> hist(time)



**Histogram of time** 

```
time
```


# **Histogram of ageelect**

> hist(yearelect)

# **Histogram of yearelect**



Linear regression:

> summary(lm(ageelect~yearelect))

Call:

 $lm(formula = ageelect ~yearelect)$ 

Coefficients:

Estimate Std. Error t value Pr(>|t|)

(Intercept) 32.726215 11.342419 2.885 0.00540 \*\*

Estimated regression equation:

ageelect =  $32.73 + 0.01812 *$  yearelect

Plotting:

reg1 <- lm(ageelect~yearelect) plot(yearelect,ageelect) abline(reg1)



# **(b)**

library(survival)

 $surv.pope = Surv(time, delta==1)$ 

fit1=survfit(surv.pope~1, type="kaplan-meier", conf.type="log-log")

#### plot(fit1)

The confidence interval here is log-log, equation (3.30) in ABG. The method for ties is the default, i.e., "kaplan-meier"



# Kaplan-Meier plot for post-election survival time

> quantile(fit1,probs=c(.25,.5,.75))

\$quantile

25 50 75

5 9 15

#### \$lower

25 50 75

2 7 12

# \$upper

25 50 75

6 11 18

```
Thus the estimates are:
Median 9 95% conf.int. (7,11)
Lower quartile 5 95% conf.int. (2,6)
Upper quartile 15 95% conf.int. (12,18)
```
**(c)**

## Consider three groups: 1400-1599, 1600-1799, 1800-today,

```
> century = cut(yearelect, breaks = c(0, 1599, 1799, 2021))
> fit2=survfit(surv.pope~century, conf.type="log-log")
> plot(fit2, main = "Kaplan-Meier plots for post-election survival 
time", lty=1:3)
```
> legend("bottomleft",c("1400-1600","1600-1800","1800-now"),lty=1:3)



# Kaplan-Meier plots for post-election survival time

#### **We perform a 3-sample logrank test:**

```
> survdiff(surv.pope~century)
Ca11:
```

```
survdiff(formula = surv.pope \sim century)
```


Chisq= 13.1 on 2 degrees of freedom, p= 0.001

We therefore reject the null hypothesis that the survival times are equally distributed for these groups.

Looking at the Kaplan-Meier plots, it might also be of interest to do a separate comparison involving only the two periods 1400-1599, 1600-1799. There may be several ways of doing this. Below we create new vectors consisting of the popes elected before 1800.

```
> newtime = time[17:63]
> newdelta = delta[17:63]
> newyear = yearelect [17:63]
> newsurv = Surv(newtime, newdelta)
> newcentury = cut (newyear, breaks = c(0, 1599, 1800))
> survdiff(newsurv~newcentury)
Call:
survdiff(formula = newsurv \sim newcentury)
                            N Observed Expected (O-E)^2/E (O-E)^2/V
newcentury=(0,1.6e+03) 28 28 22.9 1.12 2.72
newcentury=(1.6e+03,1.8e+03) 19 19 24.1 1.07 2.72
```
Chisq=  $2.7$  on 1 degrees of freedom,  $p= 0.1$ 

There is thus no significant difference between the survival in these two groups.

#### **PROBLEM 2**

```
(a)
> surv.pope = Surv(time,delta==1)
> fit.pope=coxph(surv.pope ~ ageelect + yearelect)
> summary(fit.pope)
Call:
cosh(formula = surv.pope ~ ageelect + yearelect) n= 63, number of events= 61 
              coef exp(coef) se(coef) z Pr(>|z|)ageelect 0.0428143 1.0437440 0.0139371 3.072 0.00213 ** 
yearelect -0.0034181 0.9965877 0.0008564 -3.991 6.57e-05 ***
---Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
          exp(coef) exp(-coef) lower .95 upper .95
ageelect 1.0437 0.9581 1.0156 1.0726
yearelect 0.9966 1.0034 0.9949 0.9983
Concordance= 0.678 (se = 0.041)
Likelihood ratio test= 19.72 on 2 df, p=5e-05
Wald test = 21.16 on 2 df, p=3e-05Score (logrank) test = 21.66 on 2 df, p=2e-05
```
Model:

```
alpha(t|ageelect,yaarelect) =
```

```
alpha 0(t)*exp{beta*ageeelect + beta2*yaarelect}
```
Both covariates are highly significant, and the model is significant.

The hazard ratio for ageelect is 1.0437. This means that one additional year leads to a factor of 1.0437 in the hazard.

The hazard ratio for yearelect is 0.9966. This means, for example, that 100 additional years leads to a factor of  $0.9966^{\text{A}}100 = 0.71$  in the hazard.

#### **(b)**

```
> fit.pope.inter=coxph(surv.pope ~ ageelect + yearelect + 
ageelect:yearelect)
> summary(fit.pope.inter)
```
Call:

```
cosh(formula = surv.pope ~ aqeelect + yearlect +ageelect:yearelect)
```

```
 n= 63, number of events= 61
```




```
Concordance= 0.657 (se = 0.042)
Likelihood ratio test= 21.78 on 3 df, p=7e-05
Wald test = 20.27 on 3 df, p=1e-04Score (logrank) test = 21.8 on 3 df, p=7e-05> anova(fit.pope,fit.pope.inter)
Analysis of Deviance Table
Cox model: response is surv.pope
Model 1: ~ ageelect + yearelect
Model 2: ~ ageelect + yearelect + ageelect:yearelect
    loglik Chisq Df P(>|Chi|)
1 -184.97 
2 -183.94 2.0512 1 0.1521
```
The above output says that a likelihood ratio test for the null hypothesis of no interaction has p-value 0.1521, which is not significant.

#### **(c)**

new.cov=data.frame(ageelect=c(60,60,80,80),yearelect=c(1750,1950, 1750,1950))

#### **Kaplan-Meier:**

```
surv.new=survfit(fit.pope,newdata=new.cov)
plot(surv.new,mark.time=F,lty=1:4, main="Post-election survival")
legend("bottomleft", c("60 years 1750","60 years 1950", "80 years 
1750","80 years 1950"), lty=1:4)
```
# **Post-election survival**



### **Nelson-Aalen:**

fit.na=coxph(Surv(time,delta==1)~ageelect + yearelect) surv.na=survfit(fit.na,newdata=new.cov) plot(surv.na, fun="cumhaz", mark.time=F, lty=1:4, main="Postelection survival") legend("topleft", c("60 years 1750","60 years 1950", "80 years 1750","80 years 1950"), lty=1:4)



#### **(d)**

Martingale residuals are defined in Slides 12.

We have from before:

fit.pope=coxph(surv.pope ~ ageelect + yearelect)

Now we use:

martres = fit.pope\$residuals

Residuals for ageelect:

plot(ageelect,martres)

lines(lowess(ageelect,martres))



Martingale residuals for yearelect:

plot(yearelect,martres)

lines(lowess(yearelect,martres))



The plot for yeareelct looks fine, but there may be a slight problem with ageelect (?)

### **(e)**

In order to check for possible transformations of ageelect, we fit a model with only yearelect as covariate. Then we plot the martingale residual for ageelect.

```
fit.agef=coxph(surv.pope ~ yearelect)
martres.agef = fit.agef$residuals
plot(ageelect,martres.agef)
lines(lowess(ageelect,martres.agef))
```


It might seem that the covariate ageelect shouold be transformed in some way. The behavior between 50 and 70 is OK, but the curve looks fairly horizontal up to 50 and after 70. If we ignore the part up to say 45, it might also seem that the underlying curve is logarithmic.

Below is the analysis for yearelect:

```
fit.yearf=coxph(surv.pope ~ ageelect)
martres.yearf = fit.yearf$residuals
plot(yearelect,martres.yearf)
lines(lowess(yearelect,martres.yearf))
```


This curve looks more linear, at least starting from approximately year 1500. Thus the use of year in the model looks OK.

### **Use of psplines:**

```
fit.pope1.spl=coxph(surv.pope ~pspline(ageelect,df=3) + yearelect)
termplot(fit.pope1.spl,se=T,terms=1)
```
fit.pope2.spl=coxph(surv.pope ~ageelect+pspline(yearelect,df=3)) termplot(fit.pope2.spl,se=T,terms=2)



ageelect



We chose df=3. The plots are then very close to the lowess smooths from the martingale residual plots. The conclusions are hence much the same.

## **(f)**

Schoenfeld residuals and their use are described in Slides 12.

cox.zph(fit.pope) chisq df p ageelect 0.00634 1 0.937 yearelect 3.45029 1 0.063 GLOBAL 3.74955 2 0.153

This output indicates that the beta-coefficient of ageelect does not change with t, while the one of yearelect might have a tendency to change with t. We now check this further by plotting. The beta looks rather constant for the ageelect, while there is a more strange behavior of beta for yearelect.



## **(g)**

No running of R is asked for here. We are in the situation considered on page 23/40 of Slides 12. Thus we fit a Cox-model where ageelect is the covariate, while yearelect is stratified, for example using the century variable introduced in Problem 1c:

```
fit3=survfit(coxph(surv.pope ~ ageelect + strata(century)))
```

```
plot(fit3,fun="cloglog",main="Log cumulative hazard plot for 
yearelect",lty=1:3)
```


# Log cumulative hazard plot for yearelect

We have to check whether the curves can be viewed as parallell (i.e., having a constant vertical distance). This is perhaps questionable here, confirming the possible problem with the variable as revealed by the Schoenfeld residuals.

#### **PROBLEM 3**

**Aalen's additive model**

fit.pope.aalen=aareg(surv.pope ~ ageelect + yearelect)



The tendency of these plots (looking at the "derivatives") is that things are fairly constant up to time 10, but then the derivatives increase in absolute value. This effect may also be seen in the Nelson-Aalen plots in (c ).

```
> print(fit.pope.aalen)
Call:
aareg(formula = surv.pope \sim ageelect + yearelect)
```
n= 63

27 out of 30 unique event times used

slope coef se(coef) z p Intercept 0.505000 9.28e-02 3.55e-02 2.62 0.008920 ageelect 0.013100 1.30e-03 4.51e-04 2.88 0.004040 yearelect -0.000575 -8.53e-05 2.29e-05 -3.73 0.000193

Chisq=14.67 on 2 df,  $p=0.000651$ ; test weights=Aalen

#### **PROBLEM 4**

First, restrict the dataset to popes with ageelect > 50: data.popes.50 <- subset(popes.new, ageelect>50)

time.50 = data.popes.50\$Survival

delta.50 = 1-data.popes.50\$Censored

 $ageelect.50 = data. popes.50$ \$Age.Election

yearelect.50 = data.popes.50\$Year.Elected

Then the left truncation Surv object becomes

surv.pope.left = Surv(ageelect.50,ageelect.50+time.50,delta.50==1)

A Cox-regression with "yearelect" as covariate, is then performed by

```
fit.pope.left=coxph(surv.pope.left ~)
```
summary(fit.pope.left)

new.cov.left=data.frame(yearelect.50=c(1500,1650,1800,1950))

#### **Kaplan-Meier plots:**

These are plots for the conditional distributions of age at death for popes conditional on starting their pontification at age 50 or higher.

```
surv.new.left=survfit(fit.pope.left,newdata=new.cov.left)
plot(surv.new.left,mark.time=F,lty=1:4, xlim=c(50,100),main="Popes
age at death")
legend("bottomleft", c("1500", "1650", "1800", "1950"), lty=1:4)
```
Popes age at death

